Cutting down trees with a Markov chainsaw

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- Cutting down trees
- The record point of view
- Generating random Cayley trees
- Large random Cayley trees and the CRT
- Cutting down trees to plant a forest
- Cutting down lattice paths
- Algorithms and the additive coalescent

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Natural setting: cutting at edges

- cut a random edge
- keep the bit with the root
- stop when no more edges.

Two kinds of result:

- for deterministic and log *n* trees: (weakly) 1-stable laws
- for Galton–Watson trees (\sqrt{n} trees), the case we are interested in here.

Practical setting: cutting at nodes

• stop when no more nodes

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• pick a random node

remove the subtree

History: Meir-Moon, Janson, Chassaing, Panholzer, Fill, etc... method of moments :(

Union-Find algorithms and the fragmentation/coalescent

A natural model for merge queries:

- initially have *n* sets $\{1\}, \{2\}, ..., \{n\}$
- there is an unknown spanning tree T on $1, 2, \ldots, n$
- the merge queries are the edges of T in a random order
- after n k merges the collection of sets is forest of k + 1 rooted trees \Rightarrow kernel is additive

An other important one:

- the queries are edges of K_n
- the merge queries arrive in random order
- growth of random graphs/minimum spanning tree: kernel is **multiplicative**

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Hashing with linear probing/parking



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Records I

Usual records

- take a random permutation $\pi = (\pi_1, \dots, \pi_n)$ of $\{1, 2, \dots, n\}$
- *i* is a record if $\pi_i = \min\{j \le i : \pi_j\}$.

$$\mathbf{E}\left[\sum_{i} \mathbf{1}[i \text{ is a record}]\right] = \sum_{i=1}^{n} 1/i \sim \log n$$

Equivalent to records in a rooted path P



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Records II: a tree version

Records in rooted trees: R(T) the number of records in a tree T

- for a tree *T* rooted at *r* and a random permutation π
- a vertex $u \in T$ is a record if $\pi_u = \min\{\pi_v : v \in \llbracket u, r \rrbracket\}$.

$$\mathbf{E}[R(T)] = \mathbf{E}\left[\sum_{u \in T} \mathbf{1}[u \text{ is a record}]\right] = n\mathbb{P}(U \text{ is a record}) = n\mathbf{E}\left[\frac{1}{D_U + 1}\right]$$

If T_n is a rinary search tree

$$\mathbf{E}\left[R(T_n)\right] \sim \frac{n}{2\log n}$$

If T_n is a simply generated tree

 $\mathbf{E}\left[R(T_n)\right]\approx\sqrt{n}$

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A coupling
Cut the vertices in the order given by
$$\pi$$
. Then:
 u is cut \Leftrightarrow u is a record
Cutting down trees with a Markov chainsaw
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Generating rooted labelled trees uniformly

Cayley tree:

- uniformly chosen rooted labelled tree on $\{1, 2, ..., n\}$
- uniformly random rooted spanning tree of K_n
- can also generate the shape as a Poisson(1)-Galton–Watson tree conditioned on the size being *n*.

Algorithm 1

- Start with vertex 1
- For $2 \le i \le n$ connect *i* to

$$V_i = \begin{cases} j & w.p. \ 1/n & 1 \le j \le i-2\\ i-1 & w.p. \ 1-(i-2)/n \end{cases}$$

• Randomly permute the labels

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Generating rooted labelled trees uniformly

Algorithm 2 (Aldous–Broder)

- start from a random vertex X_0 ;
- move in the graph according to a simple random walk *X*;
- if *X* visits X_i for the first time add the edge $X_{i-1} \rightarrow X_i$.
- At the cover time, return the tree of directed edges.



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Consequences for large random trees

A "building branches" point of view

Walk until the first time we go to a vertex already discovered:

- when the tree has k vertices, we stop the current branch with probability k/n.
- given that we stop, the current vertex is uniform on the already constructed tree

Lengths in Cayley trees

- $\mathbb{P}(\text{first branch} \geq k \text{ edges}) = \prod_{i=1}^{k} (1 i/n)$
- critical scaling $k \sim x\sqrt{n}$ (birthday paradox)
- \mathbb{P} (first branch $\geq x\sqrt{n}$) $\sim e^{-x^2/2}$.

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A Markov chain of spanning trees

Following the algorithm

- $\{X_j, j \ge 0\}$ a sequence of i.i.d. r.v. uniform in $\{1, 2, \dots, n\}$
- $\tau(k) = \inf\{j \ge 0 : X_j = k\}$
- *T* the tree consisting of $\{(X_{\tau(k)-1} \to X_{\tau(k)}) : \tau(k) > 0\}$



A stationary sequence of rooted trees: change the starting point

• $\{X_j, j \in \mathbb{Z}\}$ a sequence of i.i.d. r.v. uniform in $\{1, 2, \ldots, n\}$

•
$$\tau_i(k) = \inf\{j \ge i : X_j = k\}$$

• T^i the tree consisting of $\{(X_{\tau_i(k)-1} \to X_{\tau_i(k)}), \tau_i(k) > i\}$

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X_{-2}	X_{-1}	X_0	X_1	X_3	X_4	X_5	X_6	X_7
3 •	2	3	1	1	2	4	4 •	1 •

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Stick breaking construction for the limit of large trees



Scaling limit of the Aldous-Broder algorithm

- a Poisson point process with rate xdx splits $\mathbb{R}_{>0}$ into segments;
- Take the segments in order, and make them into a tree shape hooking them at a uniform position
- Taking the closure yields the continuum random tree.

A closer look at the Markov chain construction

Reformulation

- pick a random node X_{i+1}
- cut it off, together with its subtree R_{i+1}
- Build T^i by appending $T^i \setminus R_{i+1}$ as a child of X_{i+1}



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Observations about the random walk construction

Facts about the construction

Conditional on their respective sizes (subset of labels):

- left-over tree $T^1 \setminus R_1$ is a Cayley tree
- the cut-off tree R_1 is a Cayley tree

Idea:

- Run the procedure without putting edges from X_{-1} or X_0
- Conditional on $X_{-1} \neq X_0$, this is a random forest of 2 rooted trees.

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A probabilistic correspondence

Theorem

There exists a correspondence between:

- a uniform Cayley tree T and a cutting down sequence, and
- a uniform Cayley tree T^* and a uniformly random node $u \in T^*$.

In the correspondence, the length of the cutting sequence is turned into the depth of u in T^* (plus one).

Idea of the proof:

- one step of the Markov chain of spanning tree cuts at a uniformly random node;
- for the second step, rather than doing it on the whole tree, just do it on $T^0 \setminus R_1$: replace $T^0 \setminus R_1$ by the its transformation in one step of the Markov chain
- stop when you hit the root of initial tree T^0 .

A picture is worth a thousand words



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Sampling a random vertex in a tree

Algorithm

- with probability $|T|^{-1}$, return the root u_1 ;
- otherwise, pick a child u_2 of u_1 with probability proportional to $|T(u_2)|$ and recurse in $T(u_2)$

Lemma

The vertex U returned by the procedure is uniform in T.

Proof.

For all $u \in T$, there is a unique path $u_1, u_2, \ldots, u_{\kappa} = u$. Then $\mathbb{P}(U = u)$:

$$\prod_{i=1}^{\kappa-1} \left(1 - \frac{1}{|T(u_i)|}\right) \frac{|T(u_{i+1})|}{|T(u_i)| - 1} \times \frac{1}{|T(u_\kappa)|} = \frac{1}{|T|} \quad \Box$$



Lemma

At time κ , when $v_{\kappa} = X_0$, the vertex X_0 is uniformly random in $T^{\kappa-1}$

Proof.

The growth of the path $v_1, v_2, \ldots, v_{\kappa}$ follows the distribution of the path to a uniformly random node:

- with probability $|T^i(X_0)|^{-1}$ return X_0
- otherwise $T^{i+1}(X_0)$ is size-biased

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About walks associated with trees





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Let T_n be a Cayley tree of size *n*, with contour function $C(\cdot)$

Theorem (Marckert–Mokkadem) Let $e = (e(t), 0 \le t \le 1)$ be a standard Brownian excursion. Then, $\frac{C(\lfloor 2n \cdot \rfloor)}{\sqrt{n}} \xrightarrow[n \to \infty]{d} 2e(\cdot)$



Let $A = \{(x, y) : 0 \le x \le 2n \text{ and } 0 \le y \le C_T(x)\}$ Say $\pi(p) = i$ if $p \in A$ correspond to the point $u \in T$ For $p \in A$ let $\ell_p = \#\{p' : \pi(p') = \pi(p)\}$

Sampling nodes uniformly at random

- pick p = (x, y) with probability proportional to ℓ_p^{-1}
- then, $\pi(p)$ is uniform in *T*.

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Cutting down lattice paths



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2#{nodes left in the tree} = *length*{fully active region}



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Cutting down the mightiest tree with ... a Herring!

Correct scaling:

- mass by *n*
- distances by \sqrt{n}
- time by \sqrt{n}



Setting: e an excursion

- $A_e = \{p = (x, y) : 0 \le y \le e(x)\}$
- $p \in A_e$, let $\ell_p = |\{p' : \pi(p) = \pi(p')\}|$
- \mathcal{P} be a Poisson point process on $A_e \times [0,\infty)$ with intensity $\ell_{(x,y)}^{-1} dx dy dt$
- $\mathcal{P}^{\star} \subseteq \mathcal{P}$ the points landing in the active region
- $\mu(s)$ the length of the fully active region (mass left) at time s

Cutting down the mightiest tree with ... a Herring!

Poisson Point Process

Correct scaling:

- mass by *n*
- distances by \sqrt{n}
- time by \sqrt{n}



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- \mathcal{P} be a Poisson point process on $A_e \times [0, \infty)$ with intensity $\ell_{(x,y)}^{-1} dx dy dt$
- $\mathcal{P}^* \subseteq \mathcal{P}$ the points landing in the active region
- $\mu(s)$ the length of the fully active region (mass left) at time s

An other point of view

An ode to Poisson processes

- Build the CRT from the PPP on $[0, \infty)$ with intensity *xdx*
- Cut the CRT with an other PPP on $[0, \infty) \times \mathbb{R}_{>0}$ with intensity *dxdt*



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An other point of view

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- Build the CRT from the PPP on $[0, \infty)$ with intensity *xdx*
- Cut the CRT with an other PPP on $[0,\infty) \times \mathbb{R}_{\geq 0}$ with intensity *dxdt*



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Number of cuts in terms of the mass left

Lemma

Let T_n be a Cayley tree:

$$\frac{\kappa(T_n)}{\sqrt{n}} \xrightarrow[n \to \infty]{d} \int_0^\infty \mu(s) ds \stackrel{d}{=} R$$

Proof.

As $n \to \infty$

$$\frac{\kappa(T_n)}{\sqrt{n}} = \frac{1}{\sqrt{n}} \sum_{i \ge 0} \mathbf{1}[p_i \in P^\star] \approx \lim_{t \to \infty} \frac{1}{\sqrt{n}} \sum_{i=0}^{t\sqrt{n}} \mathbf{1}[p_i \in P^\star]$$
$$\approx \lim_{t \to \infty} \frac{1}{\sqrt{n}} \sum_{i=0}^{t\sqrt{n}} \frac{m(i)}{n} \approx \int_0^\infty \mu(s) ds \qquad \Box$$

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From an excursion to a reflected bridge



Theorem

In the limit, the bijection turns a Brownian excursion *e* into a Brownian bridge *B* and

$$\frac{\kappa(T_n)}{\sqrt{n}} \xrightarrow[n \to \infty]{d} L_1(B) \stackrel{d}{=} R,$$

where $(L_s(B), 0 \le s \le 1)$ is the local time process of B at 0.

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Extension to all Galton-Watson trees with finite variance

 ξ r.v. such that $\mathbf{E}\xi = 1$ and $\mathbf{Var}[\xi] = \sigma^2 < \infty$



Proof.

- at the scale (n, √n, √n) the process looks the same, but time flows more slowly by factor σ
- show that when the tree has size o(n) the number of records is $o(\sqrt{n})$

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Isolating the randomness of the cut process

Two levels of randomness

- Random tree/excursion
- Random cutting process

Theorem

Conditional on $n^{-1/2}T_n \to \mathcal{T}$ encoded by 2e,

$$\mathbb{P}\left(\left.\frac{\kappa(T_n)}{\sqrt{n}} \le x \ \middle| \ \mathcal{T}_n\right) \xrightarrow[n \to \infty]{d} \mathbb{P}\left(\left.\int_0^\infty \mu_{2e}(s)ds \le x \ \middle| \ e\right)\right.$$

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