

Wiener-Hopf Factorization: Probabilistic methods

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[summary by Jean-François Dantzer]

1. Introduction

We consider a discrete random walk on \mathbb{Z} , defined by

$$S_0 = 0 \quad \text{and} \quad S_n = \sum_{i=1}^n X_i, \quad n > 0,$$

where $(X_i)_{i \geq 1}$ is an independent identically distributed (i.i.d.) sequence of random variables. We define two hitting times ν_+ and ν_- ,

$$\nu_+ = \inf\{k > 0 / S_k > 0\}, \quad \nu_- = \inf\{k > 0 / S_k \leq 0\},$$

with the convention $\inf(\emptyset) = +\infty$.

We also define M and L two variables indicating respectively the maximum of the random walk and the hit moments at which it is attained

$$M = \sup_{n \geq 0} \{S_n\}, \quad L = \inf_{n \geq 0} \{S_n = M\}.$$

This talk presents the classical probabilistic methods to derive the joint distributions of (ν_+, S_{ν_+}) , (ν_-, S_{ν_-}) and (M, L) .

Applications of these results have been found in biology [3, 4] or in queueing theory [2]. For more details on this subject, see [1, 5].

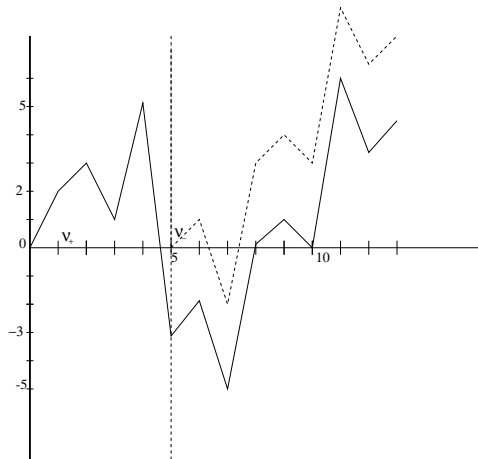


FIGURE 1. Random walks $(S_n)_{n \geq 0}$ and $(S_{n+\nu_-})_{n \geq 0}$ (in dotted line)

2. Distribution of (ν_+, S_{ν_+}) and (ν_-, S_{ν_-})

The distributions of the pairs (ν_+, S_{ν_+}) and (ν_-, S_{ν_-}) will be expressed through their generating functions, i.e., $E(u^{\nu_+} z^{S_{\nu_+}})$ and $E(u^{\nu_-} z^{S_{\nu_-}})$.

We consider three variables

$$\begin{aligned}\Psi_+(u, z) &= \frac{1}{1 - E(u^{\nu_+} z^{S_{\nu_+}})} && \text{on } \{|u| < 1, |z| \leq 1\}, \\ \Psi_-(u, z) &= \frac{1}{1 - E(u^{\nu_-} z^{S_{\nu_-}})} && \text{on } \{|u| < 1, |z| \geq 1\}, \\ \Psi(u, z) &= \sum_{n \geq 0} E(u^n z^{S_n}) && \text{on } \{|u| < 1, |z| = 1\}.\end{aligned}$$

The main result, the factorization of Wiener-Hopf described in the following proposition, gives an analytic characterization of Ψ_+ and Ψ_- .

Proposition 1. Ψ can be uniquely decomposed on $\{|z| = 1\}$ as:

$$\Psi(u, z) = \Psi_+(u, z)\Psi_-(u, z),$$

where Ψ_+ and Ψ_- have the following properties:

- Ψ_+ is analytic on $\{|z| < 1\}$;
- Ψ_+ and $1/\Psi_+$ are continuous and bounded on $\{|z| \leq 1\}$;
- $\Psi_+(u, 0) = 1$;

and

- Ψ_- is analytic on $\{|z| > 1\}$;
- Ψ_- and $1/\Psi_-$ are continuous and bounded on $\{|z| \geq 1\}$.

Proof. The proof is based on the following arguments:

- The function Ψ can be expressed with the distribution of X_1 , the sequence $(X_i)_{n \geq 1}$ is independent identically distributed, then $E(Z^{S_n}) = E(Z^{\sum_{i=1}^n X_i}) = E(Z^{X_1})^n$, thus

$$\Psi(u, z) = \sum_{n \geq 0} u^n E(z^{X_1})^n = \frac{1}{1 - uE(z^{X_1})};$$

- $(S_n)_{n \geq 0}$ and $(S_{n+\nu_-})_{n \geq 0}$ have same distribution, $(S_{n+\nu_-})_{n \geq 0}$ is the random walk beginning at the time ν_- ;
- The independence between the pair (ν_-, S_{ν_-}) and the sequence $(S_{n+\nu_-})_{n \geq 0}$;
- the same properties are also valid for $(S_{n+\nu_+})_{n \geq 0}$.

□

3. Examples

The factorization is easy when Ψ has a finite number of poles and zeros. We consider two such examples.

3.1. Random walk ± 1 . We suppose $X_i = 1$ with probability p and $X_i = -1$ with probability $(1 - p)$. In that case,

$$\Psi(u, z) = \frac{z}{-upz^2 + z - u(1 - p)}$$

Ψ has only two poles

$$\alpha_1(u) = \frac{1 - \sqrt{1 - 4u^2p(1-p)}}{2up}, \quad \alpha_2(u) = \frac{1 + \sqrt{1 - 4u^2p(1-p)}}{2up},$$

with $0 \leq \alpha_1(u) \leq 1 \leq \alpha_2(u)$.

The decomposition gives:

$$\Psi_+(u, z) = \frac{\alpha_2(u)}{\alpha_2(u) - z} \quad \text{and} \quad \Psi_-(u, z) = \frac{z}{\alpha_2(u)up(z - \alpha_1(u))}.$$

Here, we obtain the generating functions:

$$E(u^{\nu+} z^{S_{\nu+}}) = \frac{z}{\alpha_2(u)}, \quad E(u^{\nu-} z^{S_{\nu-}}) = (1 - \alpha_2(u)up) + \frac{u(1-p)}{z}.$$

3.2. Random walk left bounded. We suppose $\Pr(X_i < -1) = 0$.

$$\Psi(u, z) = \frac{1}{1 - uE(z^X)} = \frac{z}{z - uE(z^{X+1})}.$$

In that case, the factorization is easy because by Rouché's theorem the function $z \mapsto z - uE(z^{X+1})$ has one only root which belongs to $\{|z| < 1\}$, which we denote by $\alpha(u)$. One proves that $\alpha(u) \in [0, 1]$ and the decomposition of Ψ is the following:

$$\Psi_+(u, z) = \frac{z - \alpha(u)}{z - uE(z^{X+1})} \frac{u \Pr(X = -1)}{\alpha(u)}, \quad \Psi_-(u, z) = \frac{z}{z - \alpha(u)} \frac{\alpha(u)}{\Pr(X = -1)},$$

and the generating functions are:

$$E(u^{\nu-} z^{S_{\nu-}}) = 1 - \frac{u \Pr(X = -1)}{\alpha(u)} + \frac{u \Pr(X = -1)}{z},$$

$$E(u^{\nu+} z^{S_{\nu+}}) = 1 - \frac{\alpha(u)}{u \Pr(X = -1)} + \frac{z - uE(z^{X+1})}{z - \alpha(u)}.$$

4. Distribution of the Maximum and its first Hitting time (M, L)

The distribution of the pair (M, L) is expressed through its generating function $E(x^L z^M)$.

Proposition 2.

$$E(x^L z^M) = \lim_{u \rightarrow 1} \frac{\Psi_+(ux, z)}{\Psi_+(u, 1)}.$$

Proof. We define the variables M_n and L_n , as

$$M_n = \max_{0 \leq k \leq n} S_k, \quad L_n = \inf\{k / S_k = M_n\},$$

and the function H on $\{|u| < 1, |x| < 1, |z| < 1\}$ as

$$H(u, x, z) = E\left(\sum_{n \geq 0} u^n x^{L_n} z^{M_n}\right).$$

Using the same arguments as in proposition 1, it can be proved that:

$$H(u, x, z) = \frac{\Psi_+(ux, z)}{\Psi_+(u, 1)(1-u)}.$$

We conclude applying

$$\lim_{u \rightarrow 1} (1-u)H(u, x, z) = \lim_{u \rightarrow 1} (1-u)E\left(\sum_{n \geq 0} u^n x^{L_n} z^{M_n}\right) = \lim_{n \rightarrow +\infty} E(x^{L_n} z^{M_n}) = E(x^L z^M).$$

□

For the case of the random walk of subsection 3.1, it gives for $p < \frac{1}{2}$:

$$E(x^L z^M) = \frac{\alpha_2(u)}{\alpha_2(u) - z} \frac{1 - 2p}{1 - p},$$

then

$$\Pr(L < +\infty, M < +\infty) = 1,$$

and the distribution of M is geometric with parameter $\frac{p}{1-p}$, for $p > \frac{1}{2}$:

$$E(x^L z^M) = 0 \quad \text{and} \quad \Pr(L = M = +\infty) = 1.$$

References

- [1] Feller (William). – *An introduction to probability theory and its applications*. – John Wiley & Sons, New York, 1971, 2nd edition, vol. II.
- [2] Iglehart (Donald L.). – Extreme values in the $GI/G/1$ queue. *Annals of Mathematical Statistics*, vol. 43, n° 2, 1972, pp. 627–635.
- [3] Karlin (Samuel) and Altschul (Stephen F.). – Methods for assessing the statistical significance of molecular sequences features by using general scoring schemes. *Proceedings of the National Academy of Sciences of the USA*, vol. 87, 1990, pp. 2264–2268.
- [4] Karlin (Samuel) and Dembo (Amir). – Strong limit theorems of empirical functionals for large excedances of partial sums of i.i.d. variables. *Annals of Probability*, vol. 19, n° 4, 1991, pp. 1737–1755.
- [5] Spitzer (F.). – *Principles of Random Walk*. – Van Nostrand, 1964.